

3. Evolution of random graphs

Lecture 20.6.12

Notiztitel

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recall Thm 1.4:

if $p \ll \frac{1}{n}$, then $\Pr [G(n,p) \text{ has a triangle}] \xrightarrow{n \rightarrow \infty} 0$

if $p \gg \frac{1}{n}$, then $\Pr [G(n,p) \text{ has a triangle}] \xrightarrow{n \rightarrow \infty} 1$.

Definition 3.1

Let \mathcal{B} be a graph property (a class of graph closed under isomorphism).

a) A function $t(n)$ is called a **threshold function for the property \mathcal{B}** , if

$p(n) \ll t(n) \Rightarrow \Pr [G(n,p) \in \mathcal{B}] \xrightarrow{n \rightarrow \infty} 0$ „0-statement“

$p(n) \gg t(n) \Rightarrow \Pr [\quad] \xrightarrow{n \rightarrow \infty} 1$ „1-statement“

b) If one replaces $p(n) \ll t(n)$ by $\forall \varepsilon > 0 \quad p(n) \leq (1-\varepsilon) t(n)$
 $p(n) \gg t(n)$ by $\forall \varepsilon > 0 \quad p(n) \geq (1+\varepsilon) t(n)$
 and still have the same implications, then $t(n)$ is called a **sharp threshold**.

Theorem 3.2

a) $t(n) = \frac{\ln n}{n}$ is a sharp threshold for " G has no isolated vertices"

b) $t(n) = \frac{\ln n}{n}$ is a (sharp) threshold for " G is connected"

c) $\forall 0 < x < \frac{1}{2} : e^{-x-x^2} \leq 1-x \leq e^{-x}$

d) For $x \rightarrow 0 \quad 1-x \sim e^{-x}$

Proof:

c) Calculus

d) $1 \leftarrow e^{-x^2} = \frac{e^{-x-x^2}}{e^{-x}} \leq \frac{1-x}{e^{-x}} \stackrel{c)}{\leq} 1$

$$a) \quad v \in [n], \quad X_v := \begin{cases} 1 & \deg(v) = 0 \\ 0 & \text{otherwise} \end{cases}$$

$$Y := \# \text{ isol. verts} = \sum_{v \in [n]} X_v$$

$$\mu := E_X[Y] = \sum_{v \in [n]} E_X[X_v] = \sum_{v \in [n]} \Pr[X_v = 1] = n \cdot (1-p)^{n-1}$$

0-Statement: need to show: $\forall \varepsilon > 0$: If $p \leq (1-\varepsilon) \frac{\ln(n)}{n}$, then

$$\Pr[Y=0] = \Pr[G(n,p) \text{ has no isolated vertices}] \xrightarrow{n \rightarrow \infty} 0$$

$$\begin{aligned} \mu &= \frac{n}{1-p} (1-p)^n \stackrel{d!}{\sim} \frac{n}{1-p} e^{-pn} && \stackrel{d!}{\geq} \frac{n}{1-p} e^{-(1-\varepsilon)\ln(n)} \\ &= \frac{n}{1-p} n^{-(1-\varepsilon)} && = \frac{n^\varepsilon}{1-p} \xrightarrow{n \rightarrow \infty} \infty \end{aligned}$$

$$\Delta := \sum_{u \neq w} \Pr[(X_u=1) \wedge (X_w=1)] = \sum_{u \neq w} \Pr[X_u=1] \cdot \Pr[X_w=1 | X_u=1]$$

$$= \sum_{u \neq w} (1-p)^{n-1} \cdot (1-p)^{n-2} \leq n^2 (1-p)^{2n-3}$$



$$\frac{\Delta}{\mu^2} \leq \frac{n^2 (1-p)^{2n-3}}{(n (1-p)^{n-1})^2} = \frac{n^2 (1-p)^{2n-3}}{n^2 (1-p)^{2n-2}} = \frac{1}{1-p} \xrightarrow{n \rightarrow \infty} 1$$

$$\Rightarrow \Pr[Y=0] \stackrel{1.12 d)}{\leq} \underbrace{\frac{1}{\mu}}_{\rightarrow 0} + \underbrace{\frac{\Delta}{\mu^2}}_{\rightarrow 1} - 1 \xrightarrow{n \rightarrow \infty} 0.$$

1-statement need to show: $\forall \varepsilon > 0 \Rightarrow$ if $p \geq (1+\varepsilon) \frac{\ln(n)}{n}$, then

$$\Pr[Y=0] = \Pr[G(n,p) \text{ has no isolated vertices}] \xrightarrow{n \rightarrow \infty} 1$$

$\Leftarrow 1 - \Pr[Y \geq 1]$

$$\begin{aligned}
 \mu &= n(1-p)^{n-1} = \frac{n}{1-p} (1-p)^n \stackrel{c)}{\leq} \frac{n}{1-p} e^{-pn} \leq \frac{n}{1-p} e^{-(1+\varepsilon)\ln n} \\
 &= \frac{n}{1-p} n^{-(1+\varepsilon)} = \frac{n}{(1-p)n^\varepsilon} \xrightarrow{n \rightarrow \infty} 0
 \end{aligned}$$

$$\Rightarrow \Pr[Y \geq 1] \leq E_X[Y] \rightarrow 0.$$

b) Is $G(n, p)$ connected?

0-statement: For $p \leq (1-\varepsilon) \frac{\ln(n)}{n}$ we know

$$\Pr[G(n, p) \text{ connected}] \leq \Pr[G(n, p) \text{ has no isolated vertices}] \xrightarrow{a)} 0.$$

1-statement For $p \geq \frac{\ln(n)}{n}$, we want $\Pr[G(n, p) \text{ connected}] \rightarrow 1$
 $1 - \Pr[G(n, p) \text{ not connected}]$

$X_\ell := \#$ connected components of order exactly ℓ in $G(n, p)$,
 for $1 \leq \ell \leq \frac{n}{2}$.

$$Ex[X_\ell] \leq \underbrace{n^\ell}_{\# \text{ possible comp}} \underbrace{(1-p)^{\ell(n-\ell)}}_{\text{no edges to exterior}}$$



$\ell \ln(n) -$

$$\leq n^\ell e^{-p\ell(n-\ell)} = e^{\ell \ln(n) - p\ell(n-\ell)}$$

$$\leq e^{\ell \ln(n) - 6 \frac{\ln(n)}{n} \ell (n - \frac{n}{2})}$$

$$= e^{\ell \ln(n) - 3\ell \ln(n)} = e^{-2\ell \ln(n)} = n^{-2\ell} \approx \frac{1}{n^2}$$

$$\Rightarrow \sum_{\ell=1}^{n/2} Ex[X_\ell] \leq \frac{n}{2} \cdot \frac{1}{n^2} = \frac{1}{2n} \xrightarrow{n \rightarrow \infty} 0$$

$$\Rightarrow Pr[G(n, p) \text{ not connected}] = Pr\left[\sum_{\ell=1}^{n/2} X_\ell \geq 1\right] \leq Ex\left[\sum_{\ell=1}^{n/2} X_\ell\right] = \sum Ex[X_\ell] \rightarrow 0.$$

qed

Exercise 3.3

Find a function $p = p(n)$ such that for $X := \#$ spanning trees in $G(n, p)$

$$E[X] \xrightarrow{n \rightarrow \infty} \infty$$

$$Pr[X \geq 1] \xrightarrow{n \rightarrow \infty} 0$$

Theorem 3.4

Let $\rho(H) := \frac{|E(H)|}{|V(H)|}$ and call a graph balanced, iff

$\rho(H) \geq \rho(H')$ for all subgraphs $H' \subsetneq H$.

Then for a balanced graph H a threshold

for the existence of a copy of H in $G(n, p)$ is $t(n) = n^{-1/\rho(H)}$.

Proof:

Let $v := |V(H)|$ and $e := |E(H)|$. Now $t = n^{-v/e}$

Note: H, v, e are fixed, i.e. do not grow with n .

0-statement: $p \ll n^{-v/e}$. Then $p \cdot n^{v/e} = \frac{p}{n^{-v/e}} \rightarrow 0$

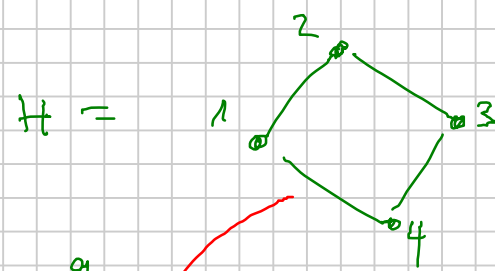
$Y := \#$ copies of H in $G(n, p)$

$$\begin{aligned} \mathbb{P}_r[H \subset G(n, p)] &= \mathbb{P}_r[Y \geq 1] \stackrel{1.10}{\leq} \mathbb{E}_x[Y] = \Theta(n^v p^e) \\ &= \Theta\left(\left(n^{\frac{v}{e}} p\right)^e\right) \longrightarrow 0. \end{aligned}$$

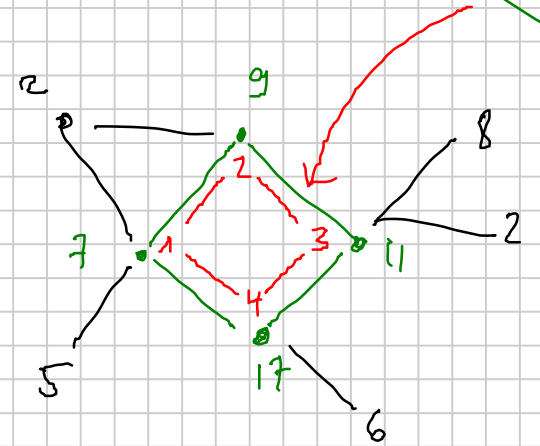
1-statement: $p \gg n^{-v/e}$. Then $p \cdot n^{v/e} = \frac{p}{n^{-v/e}} \rightarrow \infty$

Let $S \subset [n]$ be a set of size v and denote by A_S the event that $G(n, p)[S]$ contains the "lexicographic first copy" of H .

Example:

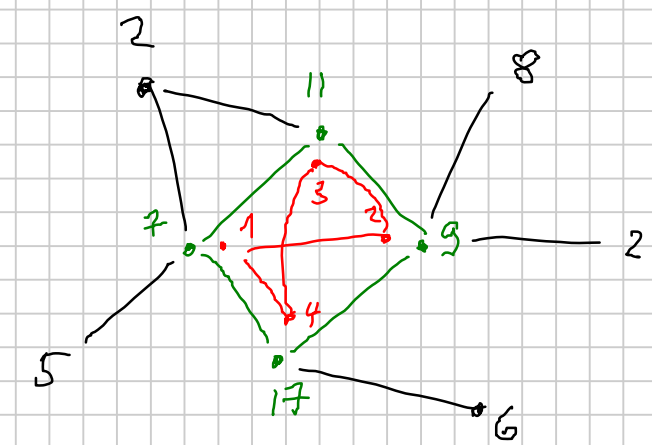


$G_1 =$



$G_1[\{7, 9, 11, 17\}]$ contains
a "lexicographic ^{first} copy" of H

$G_2 =$



$G_2[\{7, 9, 11, 17\}]$ does not
contain a lexicographic first
copy of H

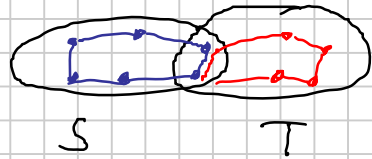
$$\text{Set } X_S := \begin{cases} 1 & A_S \\ 0 & \text{othw.} \end{cases}$$

$$Y := \sum_{S \in \binom{[n]}{v}} X_S \leq \# \text{ copies of } H \text{ in } G(n, p)$$

$$\mu := \mathbb{E}x[Y] = \sum_{S \in \binom{[n]}{v}} \mathbb{E}x[X_S] = \theta(n^v p^e) = \theta\left(\binom{n}{v} p^e\right) \rightarrow \infty.$$

Consider $S, T \in \binom{[n]}{v}$. The events A_S and A_T are independent if $|S \cap T| \leq 1$. We have to bound

$$\tilde{\Delta} := \sum_{S, T: A_S, A_T \text{ dependent}} \Pr[X_S = 1 \wedge X_T = 1] = \sum_{v'=2}^{v-1} \sum_{S, T: |S \cap T|=v'} \Pr[A_S \wedge A_T]$$



Suppose $A_S \cap A_T$ holds.

Then there are two fixed copies H_1 and H_2 of H in $G(n, p) [S \cup T]$. Denote the graph obtained by the union of H_1 and H_2 by \bar{H} , and the graph obtained by their intersection by H' .

Then $v' = |V(H')|$ and we let $e' := |E(H')|$.

Observe that $H' \subset H$, hence $\frac{e'}{v'} = \rho(H') \leq \rho(H) = \frac{e}{v}$,

thus $e' \leq v' \cdot \frac{e}{v}$.

Also, $|V(\bar{H})| = 2v - v'$ and $|E(\bar{H})| = 2e - e'$,

hence $\Pr[A_S \cap A_T] = p^{2e - e'} \leq p^{2e - v' \frac{e}{v}}$.

Summing up we obtain:

$$\begin{aligned}
 \frac{\tilde{\Delta}}{n^2} &\leq \sum_{v'=2}^{v-1} \sum_{S, T: |S \cap T|=v'} \frac{p^{2e - v' \frac{e}{v}}}{n^2} = \sum_{v'=2}^{v-1} O\left(\frac{\binom{n}{v} \binom{v}{v'} \binom{n-v}{v-v'} p^{2e - v' \frac{e}{v}}}{n^{2v} p^{2e}}\right) \\
 &= \sum_{v'=2}^{v-1} O\left(\frac{n^v n^{v-v'} p^{2e - v' \frac{e}{v}}}{n^{2v} p^{2e}}\right) = \sum_{v'=2}^{v-1} O\left(\frac{n}{n^{v'} p^{v' \frac{e}{v}}}\right) \\
 &= \sum_{v'=2}^{v-1} O\left(\left(\frac{1}{n^{\frac{v}{e}} p}\right)^{v' \frac{e}{v}}\right) \rightarrow 0.
 \end{aligned}$$

This gives

$$\Pr[\text{no copies of } H] \leq \Pr[Y=0] \leq \frac{1}{n} + \frac{\tilde{\Delta}}{n^2} \rightarrow 0.$$

qed